

On the Convergence of Credit Risk in Current Consumer Automobile Loans

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A junior credit analyst's observation

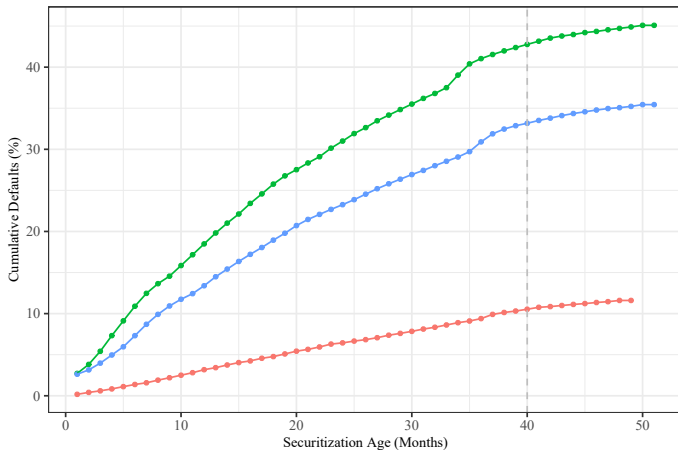


Figure: The growth rate of a cumulative ABS loss curve slows to an apparent equivalence between higher-risk or subprime pools (top curves) and lower-risk or prime pools (bottom curve).

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The Securities and Exchange Commission (SEC) recently implemented changes to the rules governing the issuance of asset-backed securities (ABS) ([Securities and Exchange Commission, 2014, 2016](#)).

Notably, it requires public issuers of ABS to make freely available pertinent loan-level information and payment performance on a monthly basis beginning in 2017.

The data in `xml` format may be accessed via the Electronic Data Gathering, Analysis, and Retrieval (EDGAR) system operated by the SEC.

Selected Bonds

We wrote Python code to scrape SEC filings to amass over 275,000 consumer automobile loans from the ABS bonds:

- ▶ CarMax Auto Owner Trust 2017-2 ([CarMax, 2017](#));
- ▶ Ally Auto Receivables Trust 2017-3 ([Ally, 2017](#));
- ▶ Santander Drive Auto Receivables Trust 2017-2 ([Santander, 2017b](#));
- ▶ Drive Auto Receivables Trust 2017-1 ([Santander, 2017a](#)).

These four bonds were selected because:

- (i) Taken together, they span the full consumer credit profile;
- (ii) No issuer is a subsidiary of an auto manufacturer;
- (iii) The paying periods span the same macroeconomic environment (i.e., actively paying starting in March-April-May 2017 for 44-52 months).

We filtered the original sample of 275,000+ loans to be as homogeneous as possible (aside from risk classification.)¹

That is, no co-borrowers, same income underwriting level (“stated not verified), no subvention, used vehicles only, loan term (72-73 months), etc.

Number of loans left for statistical analysis: **58,118**.

¹We are diversified to “noise” (i.e., largest geographic concentration: TX (13%); largest manufacturer concentration: Nissan (13%)), and we provide thorough sensitivity testing to loan selection parameters in the manuscript.

Following [Phillips \(2013\)](#), a borrower's interest rate in risk band a , r_a , is

$$r_a = r_c + m + l_a,$$

where r_c is the cost of capital, m is the added profit margin, and l_a is a factor that varies by risk band. More generally,

$$l_a \equiv f(\text{PTI}, \% \text{Down}, \text{Loan AMT}, \text{Vehicle Val.}, \text{Credit}, \text{etc.}).$$

That is, the interest rate is the market's reflection of a borrower's (i.e., loan's) risk profile.

Risk band assignment

Hence, we can defer to the market and assign borrowers to risk bands via interest rate. Specifically,

Risk Band	APR Range	Count	Default% ²
deep subprime	20%+	21,630	52%
subprime	15-20%	21,332	37%
near prime	10-15%	6,677	21%
prime	5-10%	6,300	10%
super prime	0-5%	2,179	4%
		58,118	

Note: The terms “deep subprime”, “subprime”, etc. also correspond well to the traditional credit score ranges ([Consumer Financial Protection Bureau, 2019](#)); see next slide.

²We define 3 consecutive months of missed payments = default.

Summary of 58,118 loans

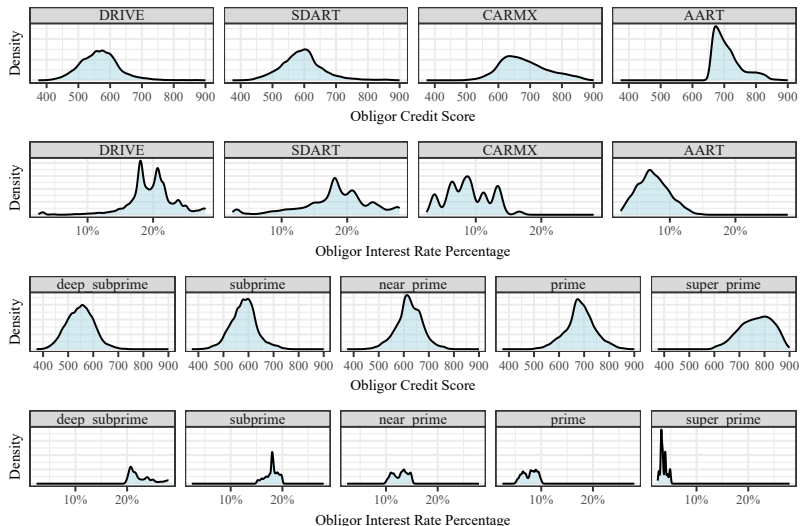


Figure: Details by bond, assigned risk band

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“Hazard”-ing a guess

Can we empirically demonstrate what appears to be an equivalence of credit risk, once a loan survives a minimal amount of time?

This is an ideal problem for the **hazard rate**,

$$\lambda(x) = \Pr(X = x \mid X \geq x),$$

because it is a *conditional* probability of default.

Goal (almost): Estimate $\lambda(x)$ by loan age and risk band for consumer loan data sampled from ABS.

We desire to differentiate between a loan contract terminating due to default or due to prepayment.

Hence, our goal is to estimate the **cause-specific hazard rate**,

$$\lambda_i(x) = \Pr(X = x, Z_x = i \mid X \geq x),$$

from loans sampled from ABS, where Z_x denotes a random variable dependent on x in the spirit of a multistate process (e.g., [Beyersmann et al., 2009](#)).

Observe $\sum_i \lambda_i(x) = \lambda(x)$.

Statistical challenges of ABS data

We are not building a model. We are letting the data completely inform an estimate of the underlying cause-specific hazard rate distribution by risk band.

When estimating the time-to-event distribution from loans sampled from a securitization pool, however, there are incomplete data challenges due to *left-truncation*, *right-censoring*, and *discrete-time*.

These challenges for ABS data are rigorously studied in [Lautier et al. \(2023a\)](#) and [Lautier et al. \(2023b\)](#).

Estimating λ_{τ}^{0i} (competing risks with censoring)

Define $f_{*,\tau}^{0i}(x) = \Pr(X = x, X \leq C, Z_x = i | X \geq Y)$, $i = 1, 2$ and $U_{\tau}(x) = \Pr(Y \leq x \leq \min(X, C) | X \geq Y)$. It may be shown

$$\lambda_{\tau}^{0i}(x) = \frac{\Pr(X = x, Z_x = i)}{\Pr(X \geq x)} = \frac{f_{*,\tau}^{0i}(x)}{U_{\tau}(x)}. \quad (1)$$

Estimation of (1) follows naturally with

$$\hat{f}_{*,\tau,n}^{0i}(x) = \frac{1}{n} \sum_{j=1}^n \mathbf{1}_{X_j \leq C_j} \mathbf{1}_{Z_{X_j} = i} \mathbf{1}_{\min(X_j, C_j) = x},$$

$$\hat{U}_{\tau,n}(x) = \frac{1}{n} \sum_{j=1}^n \mathbf{1}_{Y_j \leq x \leq \min(X_j, C_j)}$$

that is, $\hat{\lambda}_{\tau,n}^{0i}(x) = \hat{f}_{*,\tau,n}^{0i}(x) / \hat{U}_{\tau,n}(x)$, $i = 1, 2$.

Pleasingly, $\hat{\lambda}_{\tau,n}^{0i}(x)$ corresponds with classical treatments, such as [Huang and Wang \(1995\)](#), despite starting from the necessary ABS assumptions (discrete-time, sampling differences, etc.).

Hypothesis test

Lautier et al. (2023c, Proposition 1, Lemma 1), which derive the asymptotic properties of $\hat{\lambda}_{\tau,n}^{0i}$, lead to a straightforward large sample financial econometric hypothesis test.

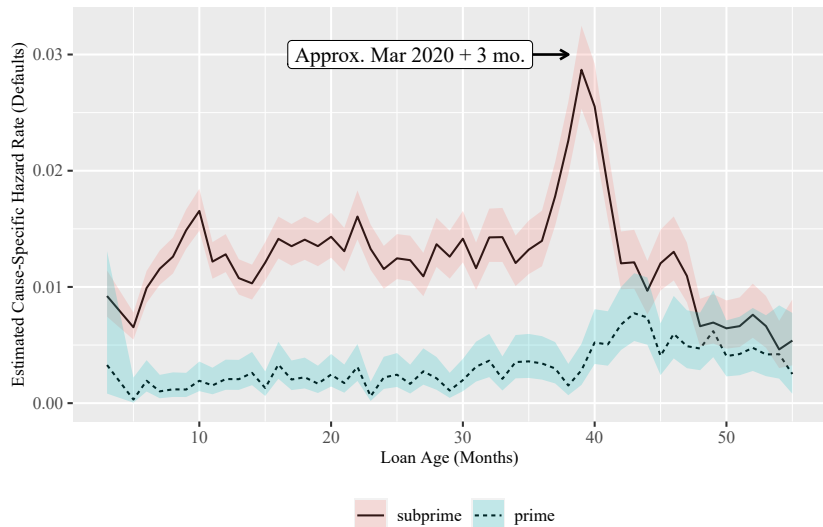
Specifically, let a, a' be two different risk bands (e.g., subprime vs. prime, etc.). Then we may test

$$H_0 : \lambda_{\tau,(a)}^{01} = \lambda_{\tau,(a')}^{01} \quad \text{vs.} \quad H_1 : \lambda_{\tau,(a)}^{01} \neq \lambda_{\tau,(a')}^{01},$$

for each age x by determining if the asymptotic confidence intervals in (2) overlap. Decision rule:

- ▶ Confidence intervals overlap \implies fail to reject $H_0 \implies$ can't claim $\lambda_{\tau,(a)}^{01} \neq \lambda_{\tau,(a')}^{01} \implies$ conditional default risk potentially converged.
- ▶ Confidence intervals **do not** overlap \implies reject $H_0 \implies$ accept $\lambda_{\tau,(a)}^{01} \neq \lambda_{\tau,(a')}^{01} \implies$ conditional default risk has **not** converged.

Credit Risk Convergence Visualization



Credit risk convergence matrix (months)

	deep sub.	subprime	near-prime	prime	sup.-prime
deep sub.	10	36	50	50	52
subprime		10	41	42	48
near-prime			10	13	43
prime				10	10
sup.-prime					10

Note: The first of three consecutive months of confidence interval overlap after month 10 for 72/73-month auto loans.

- ▶ COVID-19 plays some role but not the whole story (one-in-a-hundred-year events happen twice a decade...)
- ▶ Results hold for collateral type (new cars) and eliminating CarMax loans (different business model than banks)
- ▶ More analysis needed before generalizing to other loan types (auto loans generally considered to be a high *priority of payment* loan for consumers)

See [Lautier et al. \(2023c\)](#) for details.

Sometimes, a simple line plots suffices

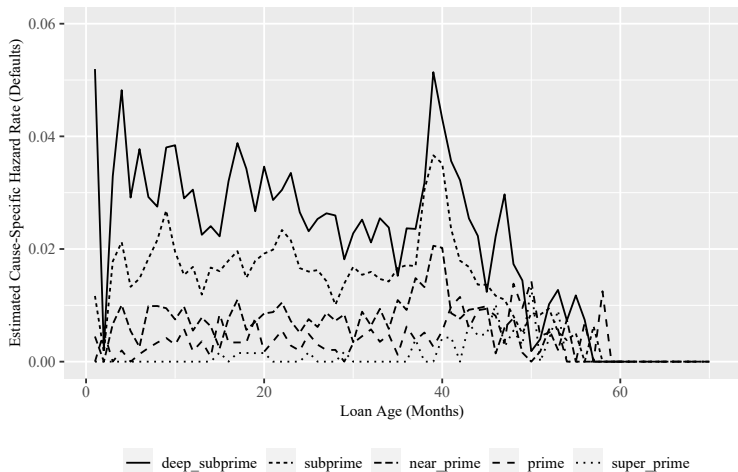


Figure: New cars, no CarMax, sans asymptotic confidence intervals

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Lender Profits “Backloaded”

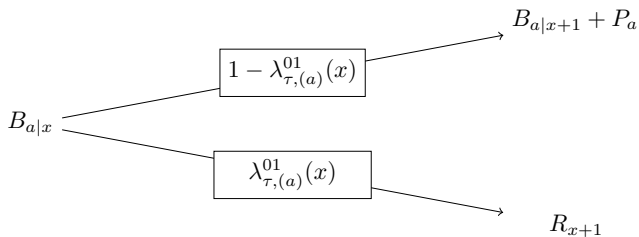
Backloaded profits necessary: Conventional wisdom is that the high-returns of high-risk loans that don't default help repay the lender for the loans that do default.

The loans we consider are sampled from securitization pools, however, and so the risk has already been transferred off the lender's books.³

Nonetheless, a risk-adjusted profitability analysis is instructive.

³ABS investors may be pricing in backloaded profits at the point of sale, but the main idea is that all our analysis is from a sample of loans the lender has already sold \implies lender portfolio level profitability not directly applicable here.

Deriving a month-by-month risk-adjusted return



Note: The assumed recovery, R_{x+1} , may be estimated from the ABS data.

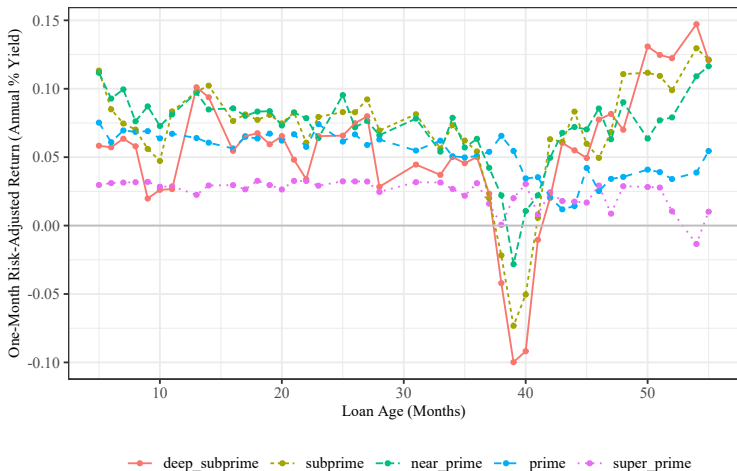


Figure: Est. conditional monthly risk-adjusted return

Estimated Savings (Deep Subprime Borrowers)

Age	#	Averages			# Pmts	Mo Pmt Savings (\$)				Total Savings (\$)				
		Bal	Pmt	APR(%)		S	NP	P	SP	S	NP	P	SP	
12	17,558	14,245	365	22.58	65									
15	16,125	13,844	364	22.56	62									
18	14,375	13,520	363	22.54	60									
24	11,628	12,836	361	22.50	56									
30	9,492	11,973	361	22.46	50									
36	7,746	10,985	359	22.46	44	16				586				
42	6,050	9,833	357	22.46	38	16				490				
48	4,899	8,799	358	22.43	33	18				438				
50	4,622	8,312	358	22.44	30	12	33	52		267	729	1,153		
54	3,568	7,485	360	22.37	26	11	30	47	61	193	531	845	1,093	
60	12	6,923	377	22.00	23	21	39	54	63	251	466	643	759	

We find that deep subprime borrowers that remain current can maximize their savings by refinancing after about 48-50 months, when they converge in risk to prime/super prime borrowers.

Most current borrowers have prepaid by about loan age 60.

Estimated Savings (Subprime Borrowers)

Age	#	Averages				# Pmts	Mo Pmt Savings (\$)				Total Savings (\$)				
		Bal	Pmt	APR(%)	S		NP	P	SP	S	NP	P	SP		
12	18,261	16,693	395	17.97	64										
15	17,021	16,126	394	17.96	61										
18	15,487	15,619	393	17.95	59										
24	12,997	14,621	389	17.94	54										
30	11,021	13,420	388	17.94	48										
36	9,309	12,194	386	17.94	42										
42	7,481	10,835	384	17.93	37		29	54			857	1,616			
48	6,192	9,506	383	17.92	31		22	44	61		526	1,055	1,473		
50	5,901	8,953	383	17.93	29		23	44	60		508	963	1,325		
54	4,542	7,975	386	17.94	25		22	40	55		389	723	988		
60	22	7,021	414	17.47	20		25	40	50		299	477	596		

We find that subprime borrowers that remain current can maximize their savings by refinancing after about 42 months, when they converge in risk to prime borrowers.

Again, most current borrowers have prepaid by about loan age 60.

Estimated Savings (Near-prime Borrowers)

Age	#	Averages				# Pmts	Mo Pmt Savings (\$)				Total Savings (\$)				
		Bal	Pmt	APR(%)	S		NP	P	SP	S	NP	P	SP		
12	5,807	19,111	411	12.79	64										
15	5,587	18,245	407	12.76	60			39						2,206	
18	5,315	17,617	405	12.74	58			40						2,158	
24	4,692	16,204	402	12.72	52			35						1,657	
30	4,146	14,694	400	12.71	47			37						1,546	
36	3,592	13,187	398	12.71	41			31						1,116	
42	3,041	11,446	394	12.67	35			28						847	
48	2,622	9,862	394	12.68	29			21	39					494	928
50	2,455	9,283	395	12.69	27			20	37					436	811
54	1,663	8,218	400	12.69	24			29	44					526	798
60	63	6,435	413	11.98	17			13	22					160	269

We find that near-prime borrowers that remain current can maximize their savings by refinancing as soon as 15 months into the loan, when they converge in risk to prime borrowers.

Surprisingly, it appears many current near-prime borrowers follow a similar prepayment pattern as deep subprime, subprime borrowers (i.e., waiting until about loan age 60).

Estimated Savings (Prime Borrowers)

Age	#	Averages				# Pmts	Mo Pmt Savings (\$)				Total Savings (\$)			
		Bal	Pmt	APR(%)	S		NP	P	SP	S	NP	P	SP	
12	5,173	18,582	358	7.83	64				39				2,327	
15	5,283	17,611	354	7.81	60				33				1,880	
18	5,315	16,706	350	7.78	57				30				1,627	
24	4,971	15,097	346	7.76	52				32				1,535	
30	4,538	13,503	345	7.74	46				30				1,245	
36	4,096	11,866	344	7.73	39				21				755	
42	3,697	10,274	342	7.72	34				23				703	
48	3,191	8,615	343	7.71	28				21				513	
50	2,963	8,101	345	7.71	26				21				460	
54	1,898	7,075	351	7.66	22				18				324	
60	92	4,756	328	7.38	16				22				261	

We find that prime borrowers that remain current can maximize their savings by refinancing as soon as 12 months into the loan, when they converge in risk to super prime borrowers.

Surprisingly, it appears many current prime borrowers follow a similar prepayment pattern as deep subprime, subprime borrowers (i.e., waiting until about loan age 60).

We can use the sibling cause-specific hazard rate estimator for prepayments, $\hat{\lambda}_{\tau,n}^{02}$, to analyze prepayment behavior by risk band.

We also overlay the Manheim Used Vehicle Value Index (ticker: MUVVI) and timing of the Economic Impact Payments for the 2017 and 2019 issuance.

Analyzing Consumer Behavior (Cont.)

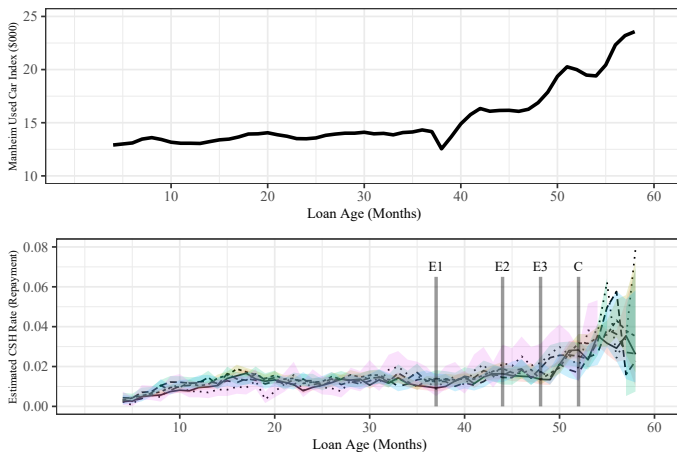


Figure: Conditional prepayment behavior by risk band

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Concluding thoughts

- ▶ *Consumerus Ignoramus?*: Consumers have a poor reputation in making financial decisions (e.g. [Gross and Souleles, 2002](#); [Stango and Zinman, 2011](#); [Lusardi and de Bassa Scheresberg, 2013](#); [Campbell, 2016](#); [Heidhues and Köszegi, 2016](#); [Dobbie et al., 2021](#)), but prepayments do accelerate as loans mature. Encourage borrowers to self-correct (questionable effectiveness (e.g., [Keys et al., 2016](#); [Agarwal et al., 2017](#))).
- ▶ *Financial innovation*: Lenders offer loans structured with a reducing payment based on good performance (may also act as an incentive to keep borrowers current keeping costs stable).
- ▶ *Competition*: Competing lenders seek out these mature loans to offer refinancing (similar to SOFI with student loans). That is, borrower delay possibly driven by perceived hassle, lack of options.
- ▶ *Regulation*: Require ongoing loans to be “re-underwritten” after a sustained period of good performance OR potentially offer borrowers cash rebates/larger trade-in values to refinance.

Thank you!

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(Lautier et al., 2023c, Proposition 1)

Define $\hat{\Lambda}_{\tau,n}^{0i} = (\hat{\lambda}_{\tau,n}^{0i}(\Delta + 1), \dots, \hat{\lambda}_{\tau,n}^{0i}(\xi))^\top$, where $\hat{\lambda}_{\tau,n}$, where $\hat{\lambda}_{\tau,n}^{0i}(x) = \hat{f}_{*,\tau,n}^{0i}(x) / \hat{U}_{\tau,n}(x)$, $i = 1, 2$. Then,

(i)

$$\hat{\Lambda}_{\tau,n}^{0i} \xrightarrow{\mathcal{P}} \Lambda_{\tau}^{0i}, \text{ as } n \rightarrow \infty;$$

(ii)

$$\sqrt{n}(\hat{\Lambda}_{\tau,n}^{0i} - \Lambda_{\tau}^{0i}) \xrightarrow{\mathcal{L}} N(0, \Sigma^{0i}), \text{ as } n \rightarrow \infty,$$

where $\Lambda_{\tau}^{0i} = (\lambda_{\tau}^{0i}(\Delta + 1), \dots, \lambda_{\tau}^{0i}(\xi))^\top$ with $\lambda_{\tau}^{0i} = f_{*,\tau}^{0i} / U_{\tau}$ and

$$\Sigma = \text{diag} \left(\frac{f_{*,\tau}^{0i}(\Delta + 1)\{U_{\tau}(\Delta + 1) - f_{*,\tau}^{0i}(\Delta + 1)\}}{U_{\tau}(\Delta + 1)^3}, \dots, \frac{f_{*,\tau}^{0i}(\xi)\{U_{\tau}(\xi) - f_{*,\tau}^{0i}(\xi)\}}{U_{\tau}(\xi)^3} \right).$$

That is, the estimators $\hat{\lambda}_{\tau,n}^{0i}(\Delta + 1), \dots, \hat{\lambda}_{\tau,n}^{0i}(\xi)$ are consistent, asymptotically normal, and independent.

Asymptotic properties (cont.)

(Lautier et al., 2023c, Lemma 1)

The $(1 - \theta)\%$ asymptotic confidence interval bounded within $(0, 1)$ for $\lambda_{\tau}^{0i}(x)$, $x \in \{\Delta + 1, \dots, \xi\}$, $i = 1, 2$ is

$$\exp \left\{ \ln \hat{\lambda}_{\tau,n}^{0i}(x) \pm \mathcal{Z}_{(1-\theta/2)} \sqrt{\frac{\hat{U}_{\tau,n}(x) - \hat{f}_{*,\tau,n}^{0i}(x)}{n \hat{U}_{\tau,n}(x) \hat{f}_{*,\tau,n}^{0i}(x)}}} \right\}, \quad (2)$$

where $\mathcal{Z}_{(1-\theta/2)}$ represents the $(1 - \theta/2)$ th percentile of the standard normal distribution.

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